

Clustering in multivariate extremes

A case study of Dutch hydrological extremes

Lambert De Monte¹

Joint work with Bas Matsuura² and Ioannis Papastathopoulos¹



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School of Mathematics

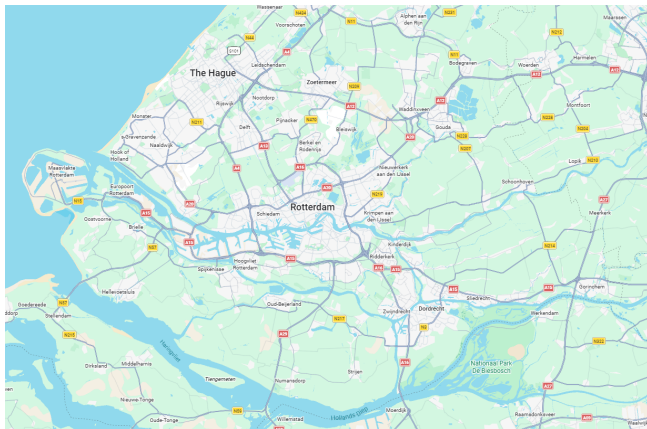


Koninklijk Nederlands
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Ministerie van Infrastructuur en Waterstaat

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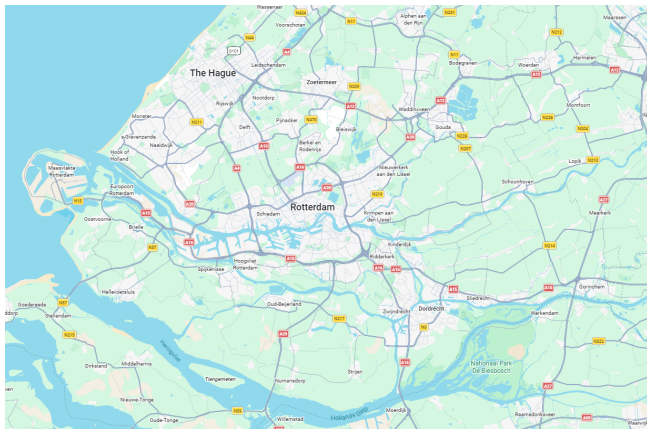
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- The Netherlands is vulnerable to compound hydrological extremes (Rhine, Meuse, North Sea).
- Extremes persist and cluster across days; duration matters for risk and impact.
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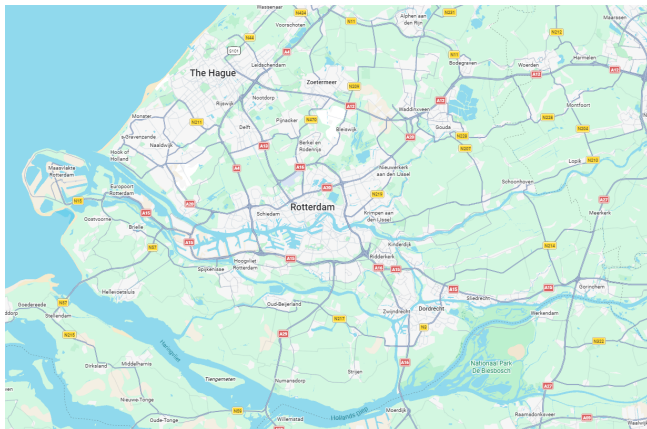
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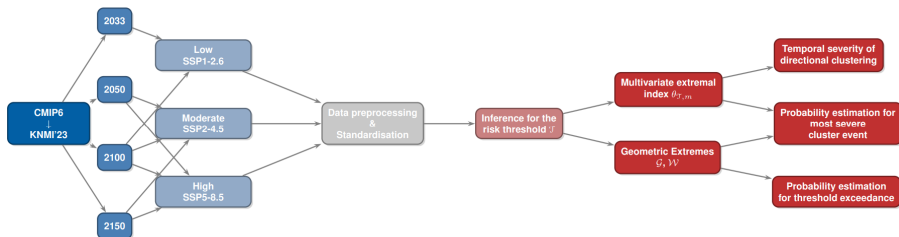


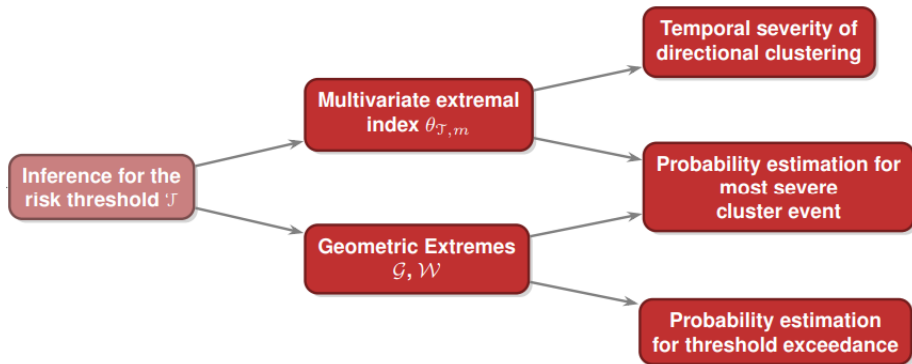
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Modelling workflow





Extremal index (univariate): Block maxima, POT, and Runs estimator

- Consider a sequence $\{X_i\}_{i=1}^n$ and define the block maximum $M_n = \max(X_1, \dots, X_n)$.

- IID case (in the MDA of GEV):

$$\mathbb{P}\left(\frac{M_n - b_n}{a_n} \leq x\right) \longrightarrow G(x; \xi) = \exp\left\{- (1 + \xi x)^{-1/\xi}\right\}, \quad 1 + \xi x > 0.$$

- With temporal dependence, *i.e.* stationary sequences, and extremal index $\theta \in (0, 1]$:

$$P\left(\frac{M_n - b_n}{a_n} \leq x\right) \longrightarrow G_\xi(x)^\theta.$$

- Consequences: adjust POT inference for dependence via declustering or θ ; mean cluster size $\approx 1/\theta$.
- Runs estimator (O'Brien 1987):

$$\theta_{u,m} = \mathbb{P}(X_{t+1} \leq u, \dots, X_{t+m} \leq u \mid X_t > u)$$

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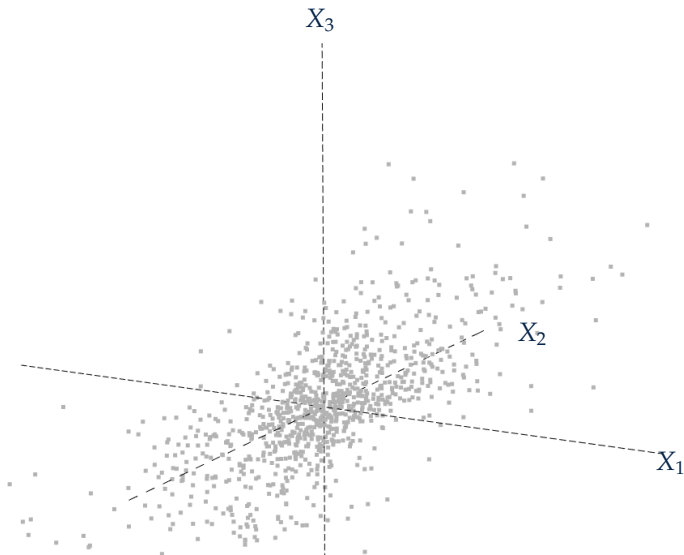
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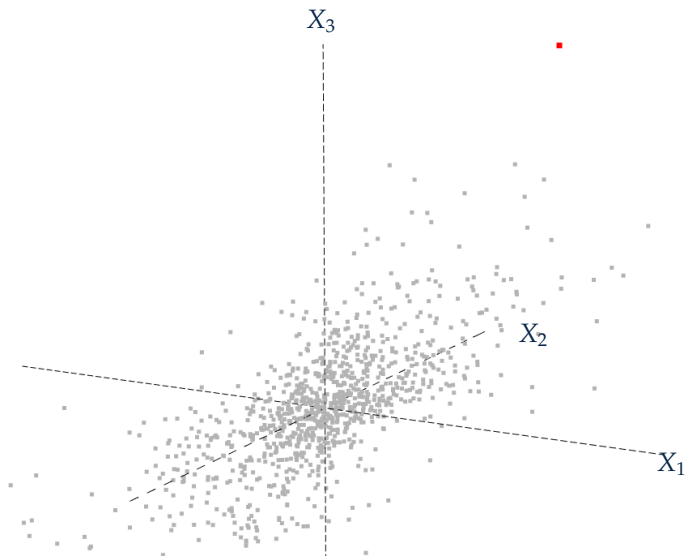
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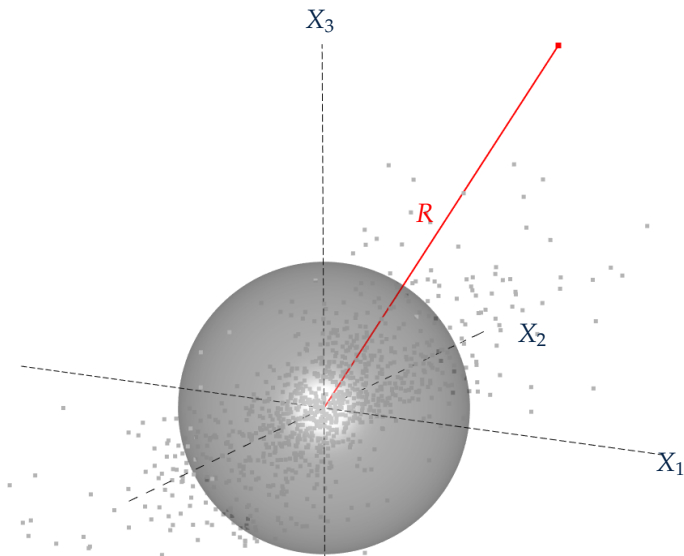
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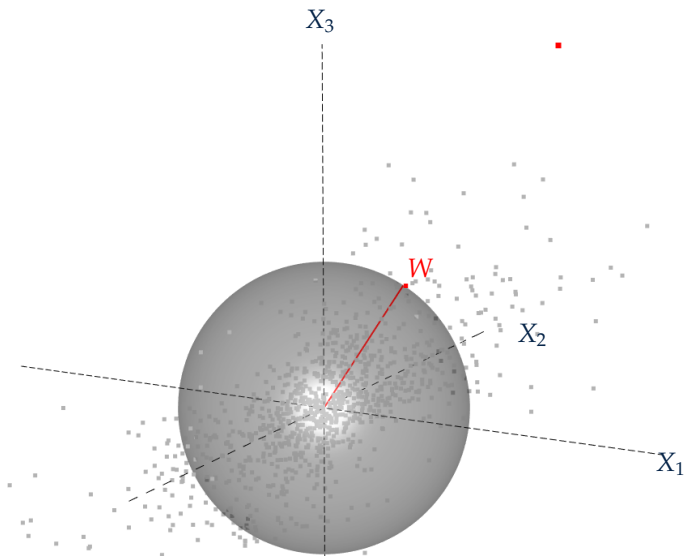
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Definition (Multivariate threshold)

A set $\mathcal{T} \subset \mathbb{R}^d$ with the origin $\mathbf{0} \in \mathcal{T}$ defines a valid radial-directional threshold via its boundary $\partial\mathcal{T}$ if $\partial\mathcal{T}$ can be represented as a function $\partial\mathcal{T} : \mathbb{S}^{d-1} \rightarrow [0, \infty)$ and if

$$\mathcal{T} := \{rw \in \mathbb{R}^d : 0 \leq r \leq \partial\mathcal{T}(w)\},$$

for all $w \in \mathbb{S}^{d-1}$.

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for all $\mathbf{w} \in \mathbb{S}^{d-1}$.

Definition (Fixed radial quantile threshold)

A set $\mathcal{T} \subset \mathbb{R}^d$ is a quantile set of probability content $q \in (0, 1)$ if its boundary $\partial\mathcal{T}$ is defined by

$$\partial\mathcal{T}(\mathbf{w}) = F_{R|\mathbf{W}}^{\leftarrow}(q \mid \mathbf{w}) = \inf\{r > 0 : F_{R|\mathbf{W}}(r \mid \mathbf{w}) \geq q\},$$

where $F_{R|\mathbf{W}}$ and $F_{R|\mathbf{W}}^{\leftarrow}$ denote the distribution and quantile functions of $R \mid \mathbf{W}$, respectively.

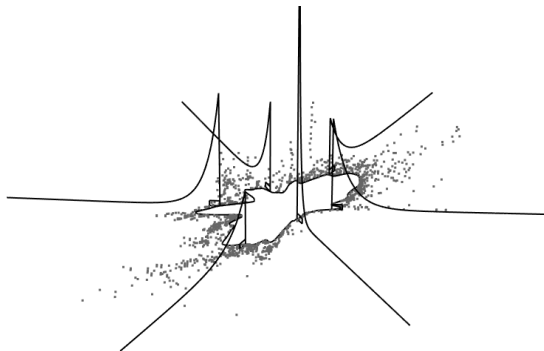
Recent radial-directional approaches

Each of these approaches (in footnote) rely on something like

$$R \mid \{R \text{ is large enough}, \mathbf{W} = \mathbf{w}\} \sim \text{SomeUnivariateModel}$$

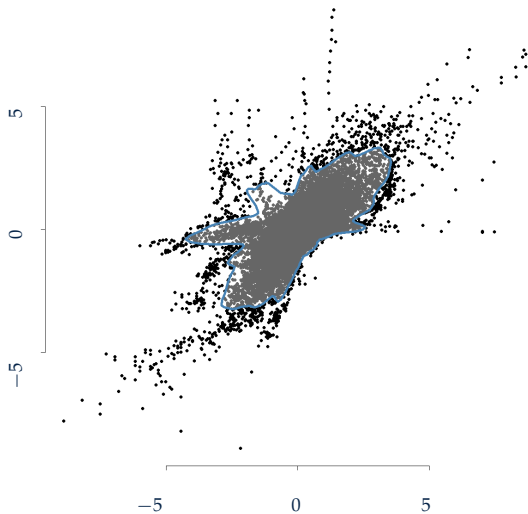
or

$$R - \partial\mathcal{T}(\mathbf{W}) \mid \{R > \partial\mathcal{T}(\mathbf{W}), \mathbf{W} = \mathbf{w}\} \sim \text{SomeUnivariateModel}$$



Wadsworth & Campbell (2024), Simpson & Tawn (2024), Papastathopoulos et al. (2025), Murphy-Barltrop et al. (2024), De Monte et al. (2025), Murphy-Barltrop et al. (2025)

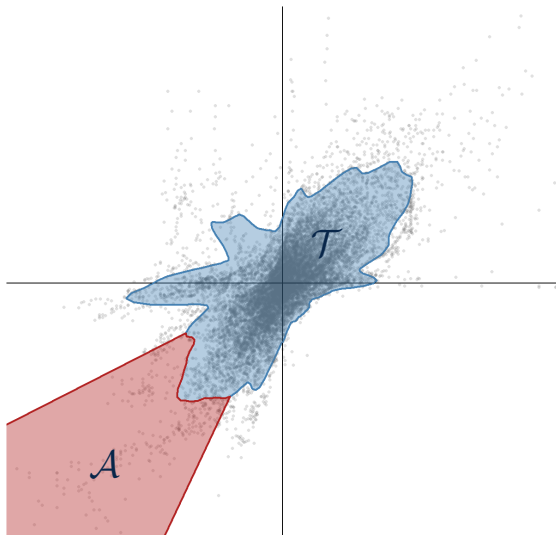
Temporal clustering of extreme events



Multivariate extremal index

For some $\mathcal{A} \subseteq \mathbb{S}^{d-1}$, we define the directional extremal index

$$\Theta_{\mathcal{T},m}(\mathcal{A}) := \mathbb{P} \left[X_{i+1} \in \mathcal{T}, \dots, X_{i+m} \in \mathcal{T} \mid X_i \notin \mathcal{T}, \frac{X_i}{\|X_i\|} \in \mathcal{A} \right]$$



Direction-wise path of exceedances

Directional extremal index

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Definition

A cluster $\mathcal{C}$ of exceedances of a threshold $\mathcal{T}$ of runs length m is the sequence of all successive times between and including some $i_j, i_{j'} \in \mathcal{I}$ with $j \leq j'$ such that $(i_j - i_{j-1}) > m$, $(i_{j'+1} - i_{j'}) > m$, and $i_k - i_{k+1} \leq m$ for any $j \leq k < j'$. This cluster has size $\#\mathcal{C} = (i_{j'} - i_j + 1)$.

Directional mean cluster size (runs of length m):

$$\begin{aligned} C_{T,m}(w) &:= \frac{1}{\theta_{T,m}(w)} \\ &= \mathbb{E} \left[\text{"Number of exceedances before } m \text{ non-exceedances"} \mid X \notin \mathcal{T}, \frac{X}{\|X\|} = w \right] \end{aligned}$$

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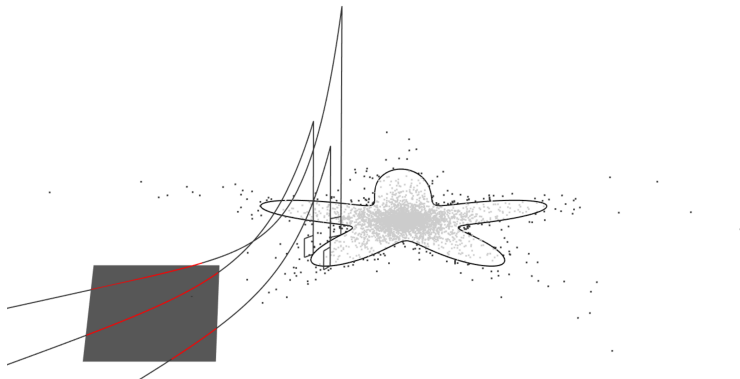
Aggregated mean cluster size (runs of length m):

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Probability estimation – A cluster's radial maxima

The probability that the radial-maximum event in a cluster lands in a risk region $\mathcal{R}$ is given by

$$\mathbb{P}[c_{\text{RM}}(\mathcal{C}) \in \mathcal{R}] = (1 - q) \int_{\mathbb{S}^{d-1}} \int_{\partial\mathcal{T}(w)}^{\infty} \mathbb{1}_{rw}[\mathcal{R}] \theta(w) \exp\left[\frac{r - \partial\mathcal{T}(w)}{\Sigma(w)}\right] f_W(w) \mathrm{d}r \mathrm{d}w.$$



1) Standardise and learn a multivariate threshold

- Margins transformed to standard Laplace for geometric EVT.
- Infer $\mathcal{T}$ (e.g., quantile set) from data with uncertainty.
- SPDE/GMRF on $\mathbb{S}^{d-1}$ for flexible direction-dependent radial functions.

2) Learn the directional extremal index $\theta_{T,m}$

- Build binary outcomes after exceedances: $z_i = 1$ if next m observations fall in T , else 0.
- Logistic link over directions: $\theta_{T,m}(w) = \frac{1}{1 + \exp\{\beta + s(w)\}}$ with $s(w)$ a cyclic/spherical spline with GMRF prior.
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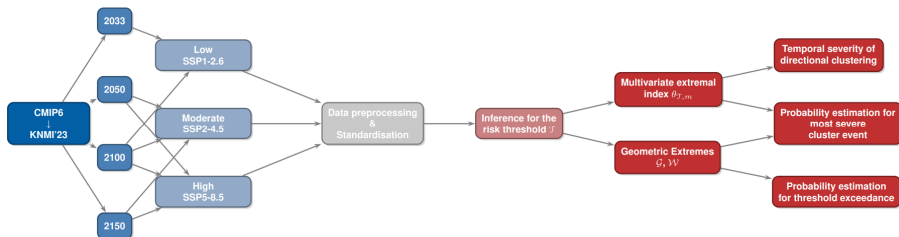
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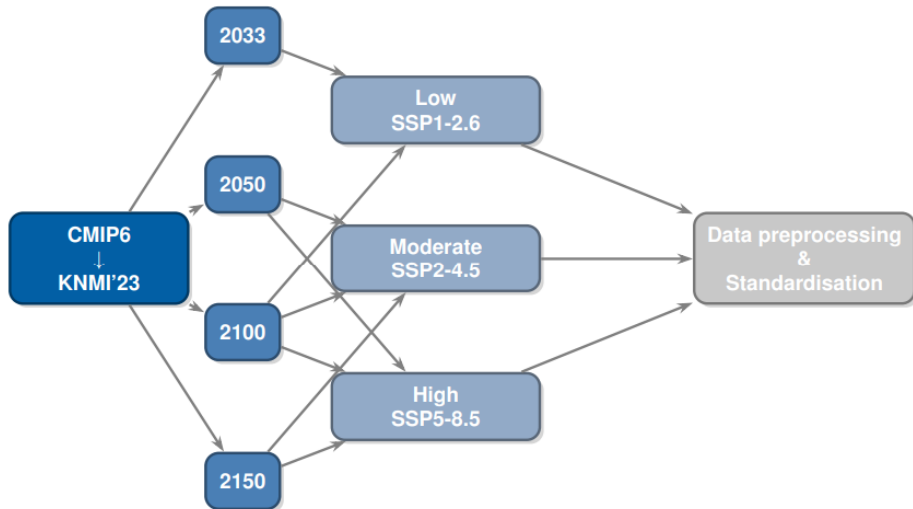
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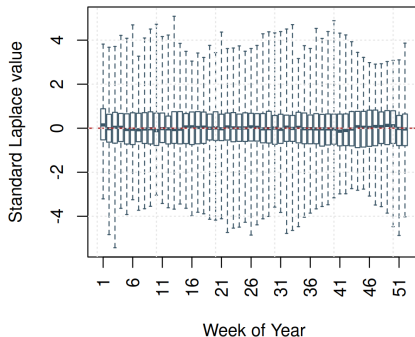
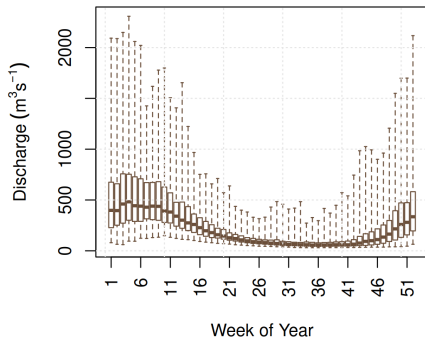
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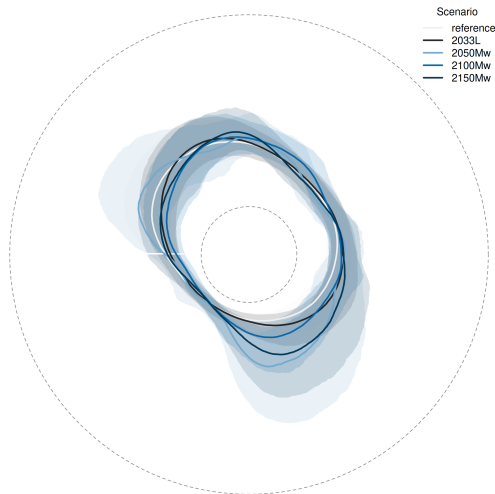


Homogenisation



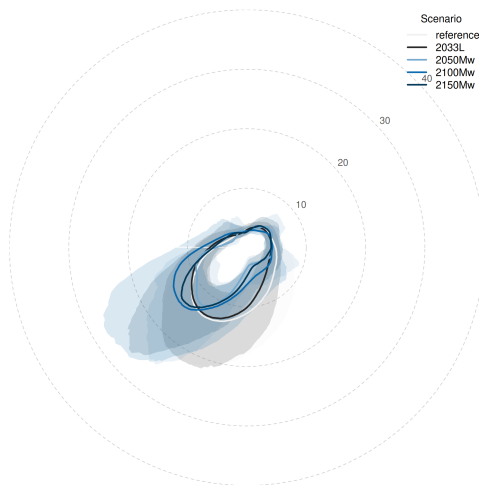
Boxplots of the yearly- and weekly-aggregated original and homogenized discharges for the Meuse and Rhine, pooling all samples across ensemble members from the 2033L scenario.

Directional extremal index



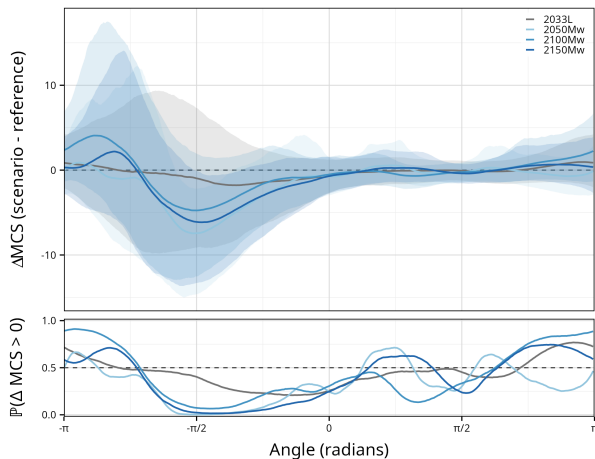
Posterior mean and 95% marginal credible intervals of the extremal index for the scenarios reference, 2033L, 2050Mw, 2100Mw and 2150Mw based on $q = 0.99$ and a runs declustering length of $m = 10$.

Directional mean cluster size



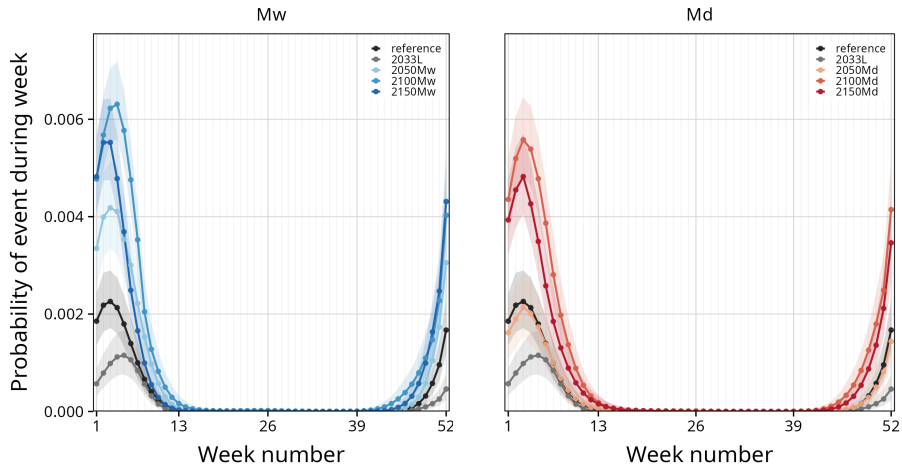
Posterior mean and 95% marginal credible intervals of the mean cluster size for the scenarios reference, 2033L, 2050Mw, 2100Mw and 2150Mw based on $q = 0.99$ and a runs declustering length of $m = 10$.

Direction-wise clustering comparison



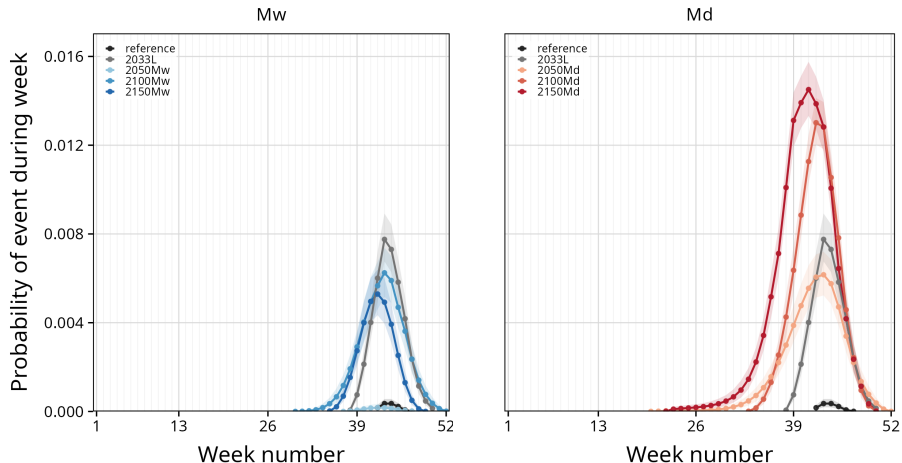
Posterior comparison of directional mean cluster size relative to the reference scenario for the Mw pathway based on $q = 0.99$ and runs declustering length $m = 10$. The upper panels show the posterior median difference $\Delta\text{MCS}(\mathbf{w}) = \text{MCS}_s(\mathbf{w}) - \text{MCS}_{\text{ref}}(\mathbf{w})$ with 95% credible intervals as a function of angular direction $\mathbf{w}$. The lower panels show the posterior probability $\mathbb{P}(\Delta\text{MCS}(\mathbf{w}) > 0)$, with the dashed line indicating probability 0.5.

Probability of high-surge events



Weekly probability estimates of Mw (left) and Md (right) scenarios across time horizons for $\mathbb{P}(Q_R \geq 12,000, Q_M \geq 3,000)$.

Probability of low-surge events



Weekly probability estimates of Mw (left) and Md (right) scenarios across time horizons for $\mathbb{P}(Q_R \leq 500, Q_M \leq 20)$.

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 - Adapt runs-based declustering methods from univariate theory to form approximately independent cluster.
 - Cluster functional: radial-maximum (analogue of univariate cluster max).
 - Estimation of direction-specific mean cluster sizes $C_{T,m}(w)$.
 - Posterior predictive probabilities for cluster-level events (return regions/periods).
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 - First time a return period is applied to a hydrological variable with a spatially varying return period.
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 - First time a significant impact is quantified on the frequency of extreme sea levels.
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 - Posterior predictive probabilities for cluster-level events (return regions/periods).
- Hydrological extremes near Rotterdam:
 - Little evidence against a change of clustering with respect to “new normals”.
 - Most of the change in the hydrological extremes appear to be marginal.
 - Relatively clear change in the frequency of extremes of low and high discharges with climate change.

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