

Non-Parametric Simulation of Multivariate Extreme Events via Spectral Bootstrap

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Elements of extreme value theory

Goals of Extreme Value Theory



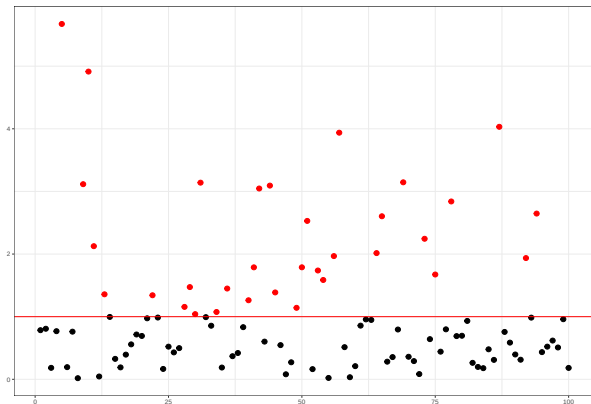
Goals of Extreme Value Theory

1. Estimate the probability of occurrence of an event more severe/extreme than previously observed
2. Estimate an extreme quantile

⇒ Inference outside the sample support

Univariate Peaks-over-Threshold method

- $Y_1, Y_2, \dots$ a series of i.i.d. random variables
- Fix a (high) threshold u
- **Extreme event** = Y_i exceeds u
 - Given that $Y_i > u$, an **excess** is defined by $Z_i = Y_i - u$



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- Fix a (high) threshold u
- **Extreme event** = Y_i exceeds u
→ Given that $Y_i > u$, an **excess** is defined by $Z_i = Y_i - u$
- Excess distribution

$$\bar{F}_u(z) = P[Y_1 - u > z \mid Y_1 > u] = \frac{\bar{F}(u+z)}{\bar{F}(u)}, \quad z > 0.$$

Balkema et de Haan (1974), Pickands (1975)

Under certain conditions, the distribution of excesses F_u converges, as $u \rightarrow \infty$, to a generalized Pareto distribution (GPD) whose distribution function is

$$H_{\sigma, \gamma}(z) = \begin{cases} 1 - \left(1 + \frac{\gamma}{\sigma} z\right)^{-1/\gamma} & \text{if } \gamma \neq 0 \\ 1 - \exp\left(-\frac{z}{\sigma}\right) & \text{if } \gamma = 0 \end{cases}$$

- Families of possible distributions for excesses = parametric family
→ **Generalized Pareto distributions (GPD)**

Generalized Pareto distributions

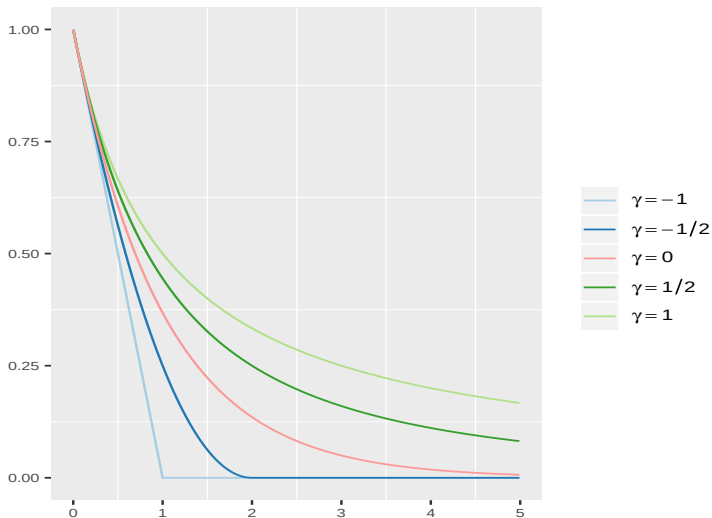


Figure: GPD survival functions

3 domains of attraction

1. Fréchet domain ($\gamma > 0$): **heavy-tailed distributions**

$$1 - H_\gamma(z) \underset{+\infty}{\sim} \gamma^{-1/\gamma} z^{-1/\gamma}$$

Examples: Cauchy, Log-gamma, Student

2. Gumbel domain ($\gamma = 0$): **thin tail distributions**

$$1 - H_0(z) \underset{+\infty}{\sim} \exp(-z)$$

Examples: Gaussian, Gamma, Exponential

3. Weibull domain ($\gamma < 0$): **finite tail distributions**

$$1 - H_\gamma(z) = 0 \quad \text{for } x \geq -1/\gamma$$

Examples: Uniform, Beta

Non parametric simulation of multivariate extremes via spectral bootstrap

Motivation and goal

- Extreme Value Theory (EVT) focuses on behaviour in the tail of multivariate distributions.
- Few extreme observations → difficult inference
→ Natural idea = bootstrap to increase the number of data
- BUT, standard bootstrap fails for EVT

Goal

Non-parametric simulation approach to expand the number of observations above the extreme level

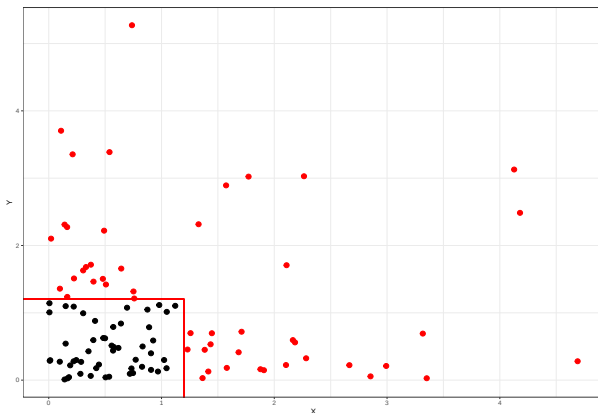
→ Extension of the non-parametric simulation algorithm for multivariate extremes, developed by Legrand et al. [2023] in the dimension 2 case to larger dimensions

Main contributions

- Expands the number of observation above extreme level
 - Ensures more reliable estimations
 - Enables extrapolation beyond the range of observed data
- Makes use of a **spectral representation** of multivariate extreme vectors

Multivariate Generalized Pareto Distributions

- $\mathbf{X} = (X_1, \dots, X_d)$ observations
- Choose (high) thresholds $\mathbf{u} = (u_1, \dots, u_d)$
- **Extreme event** = AT LEAST one of the X_j exceeds its threshold u_j



Multivariate Generalized Pareto Distributions

- $\mathbf{X} = (X_1, \dots, X_d)$ observations
- Choose (high) thresholds $\mathbf{u} = (u_1, \dots, u_d)$
- **Extreme event** = AT LEAST one of the X_j exceeds its threshold u_j
- Theory: asymptotically (when $\mathbf{u} \rightarrow \infty$), exceedances occur according to a Poisson process and $\mathbf{Z} = \mathbf{X} - \mathbf{u} \mid \mathbf{X} \not\leq \mathbf{u}$ follows a **multivariate Generalized Pareto Distribution (MGPD)** with a scale parameter σ and a shape parameter γ .
- Standard MGPD $\rightarrow \sigma = \mathbf{1}$ and $\gamma = \mathbf{0}$
 $\rightarrow$ Exponential marginals



NO parametric family of limits distributions

- Rootzén et al. [2018], Kiriliouk et al. [2019] have proposed explicit density formula for specific models
- Non parametric models are difficult to fit (lack of data)

Multivariate generalized Pareto Vectors

- Let $\mathbf{X}$ be a d -dimensional random vector, the vector of excesses is defined as

$$\mathbf{Z} = \mathbf{X} - \mathbf{u} \mid \mathbf{X} \not\leq \mathbf{u} \quad (1)$$

where $\mathbf{u} \in \mathbb{R}^d$ is a vector of suitably chosen thresholds and $\not\leq$ means that at least one of the components of $\mathbf{X} - \mathbf{u}$ is positive.

- Rootzén et al. [2018] have shown that a standard MGP vector $\mathbf{Z}$ can be decomposed as follows

$$\mathbf{Z} = E + \mathbf{T} - \max(\mathbf{T}),$$

where

- E is a unit exponential variable ;
- $\mathbf{T}$ a d -dimensional random vector
- $\mathbf{T}$ and E are independent.

Non-parametric joint MGP simulation in dimension 2

Legrand et al. [2023]

- From Rootzén et al. [2018b],

$$\begin{cases} Z_1 &= E + T_1 - \max(T_1, T_2) \\ Z_2 &= E + T_2 - \max(T_1, T_2) \end{cases}$$

- Noting $\Delta = Z_1 - Z_2 = T_1 - T_2$,

$$\begin{cases} Z_1 &= E + \Delta \mathbf{1}_{\Delta < 0} \\ Z_2 &= E - \Delta \mathbf{1}_{\Delta \geq 0} \end{cases}$$

- Simulate values of Δ and E independently
- Simulate E is trivial
- Difficulty : simulate Δ
 $\Rightarrow$ **Bootstrapping on Δ**

Spectral Representation of MGP en dimension d

Key theorem (Rootzén et al., 2018)

A standard MGP vector $\mathbf{Z}$ admits representation:

$$\mathbf{Z} = \mathbf{E} + \mathbf{T} - \max_k T_k,$$

where $E \sim \text{Exp}(1)$ and $\mathbf{T}$ is a spectral random vector.

- Define differences:

$$\Delta_j = Z_j - \max_k Z_k.$$

- Then $\mathbf{Z} = \mathbf{E} + \mathbf{\Delta}$.
- Spectral bootstrap uses bootstrap on observed $\mathbf{\Delta}$ values.

Algorithm 1: Multivariate Extreme Spectral Bootstrap

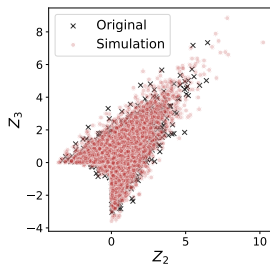
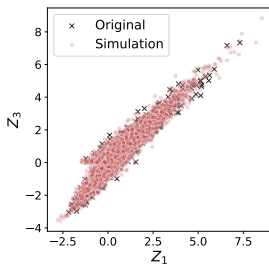
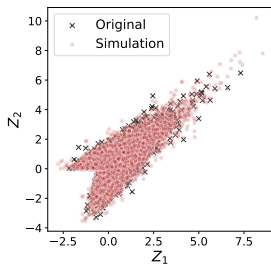
Algorithm

- 1: **require** n i.i.d. observations $\mathbf{Z}_1, \dots, \mathbf{Z}_n$ of a standard MGP distribution in $\mathbb{R}^d$
- 2: **output** m standard MGP simulated samples $\mathbf{Z}_1^*, \dots, \mathbf{Z}_m^*$ in $\mathbb{R}^d$
- 3: **procedure**
- 4: Compute $\Delta_{i,j} \leftarrow Z_{i,j} - \max_{1 \leq k \leq d} Z_{i,k}$, for $1 \leq i \leq n$, $1 \leq j \leq d$
- 5: Generate $E_1, \dots, E_m \stackrel{\text{i.i.d.}}{\sim} \text{Exp}(1)$, independently of $\Delta_1, \dots, \Delta_n$
- 6: where $\Delta_i = (\Delta_{i,1}, \dots, \Delta_{i,d})$ for $1 \leq i \leq n$
- 7: Generate m bootstrap samples $\Delta_1^*, \dots, \Delta_m^*$ from $(\Delta_1, \dots, \Delta_n)$
- 8: **end procedure**
- 9: **return** $\mathbf{Z}_\ell^* \leftarrow E_\ell + \Delta_\ell^*$ for $1 \leq \ell \leq m$

- Produces synthetic MGP observations preserving tail dependence.
- Theoretical guarantees that the samples obtained with Algorithm ?? are distributed according to the same distribution as the original observations.
- The SB simulation procedure is applied to standard MGP vectors so that the procedure itself ensures the preservation of the dependence structure of the original MGP vectors.

Non-parametric joint MGP simulation

Bivariate representations of the simulated sample of size 2,000 of a MGP vector $\mathbf{Z} \in \mathbb{R}^3$ with $\mathbf{T}$



Non-parametric joint MGP simulation

Bivariate representations of the simulated sample of size 2,000 of a MGP vector $\mathbf{Z} \in \mathbb{R}^3$ with T

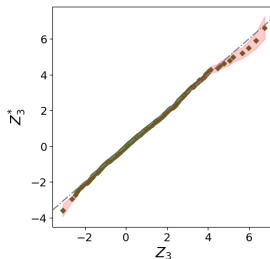
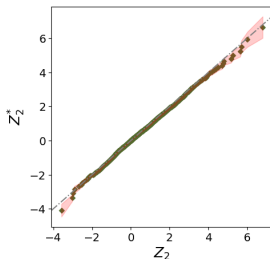
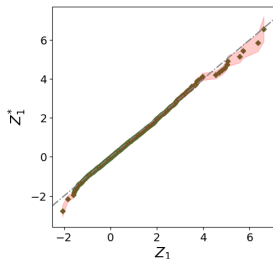


Illustration on Risk management

Illustration on Risk management

- **Risk management** = Crucial in various sectors, specific to each sector
- **Risk factor** = any variable that could result in a loss or damage
 - can be represented by random variables that quantify the magnitude of potential losses
- **Examples**
 - Climatology
 - Meteorological and marine hazards can cause significant damages
 - e.g., drought, floods, landslides
 - Risk factor = any physical quantity (wave heights, wind gusts, precipitation)
 - Finance
 - Market movements can result to substantial losses
 - Risk factor = typically, market parameters, interest rates or exchange rates
- **Tail risk:** Events with very severe magnitudes and that occurred with very low probability

Aim

For a given target risk factor X_j , accurately quantify its risk through the estimation of tail risk metrics (TRM).

Tail risk metrics

- $\mathbf{X} = (X_1, \dots, X_d) \in \mathbb{R}^d$ vector of risk factors, with X_j of density f_j
- We consider 3 TRMs defined as
 - **Expected Shortfall** [Artzner et al., 1999] at level $\alpha \in (0, 1)$

$$\text{ES}_\alpha(X_j) = \mathbb{E}[X_j \mid X_j > \text{VaR}_\alpha(X_j)] = \frac{1}{1 - \alpha} \int_{\text{VaR}_\alpha(X_j)}^{\infty} x_j f(x_j) dx_j$$

where $\text{VaR}_\alpha(X_j) := \inf\{x_j \in \mathbb{R} : \mathbb{P}(X_j \leq x_j) \geq 1 - \alpha\}$.

- **Multivariate marginal ES** of X_j at level α

$$\text{MMES}_\alpha(X_j; \mathbf{X}) = \mathbb{E}[X_j \mid \mathbf{X}_{-j} \geq \mathbf{v}_{-j}^\alpha] = \int_{\mathbb{R}} x_j f_{X_j \mid \mathbf{X}_{-j} \geq \mathbf{v}_{-j}^\alpha}(x_j) dx_j,$$

with $\mathbf{v}^\alpha = \text{VaR}_\alpha(\mathbf{X}) \in \mathbb{R}^d$

- **Dependent conditional tail expectation** of X_j at level α

$$\text{DCTE}_\alpha(X_j; \mathbf{X}) = \mathbb{E}[X_j \mid \mathbf{X} \geq \mathbf{v}^\alpha] = \int_{v_j^\alpha}^{\infty} x_j f_{X_j \mid \mathbf{X} \geq \mathbf{v}^\alpha}(x_j) dx_j,$$

Simulation framework

- Let $\mathbf{X} = (X_1, X_2, X_3)$ be a random vector with marginals distributed as a Student t -distribution with degrees of freedom $\nu_1 = 2$, $\nu_2 = 3$, $\nu_3 = 2.5$.
- An underlying assumption of our simulation framework is that the components of $\mathbf{X}$ are asymptotically dependent
- To ensure that this hypothesis is satisfied, we consider the Gumbel copula [Nelsen, 2006] to obtain dependent extremes in the upper tail

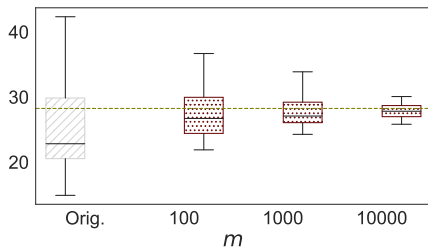
$$\mathcal{C}(\mathbf{y}) := \exp \left(- \left(\sum_{i=1}^3 [-\log(y_i)]^\theta \right)^{1/\theta} \right),$$

where $\theta \geq 1$ is the copula parameter. The larger θ , the stronger the asymptotic dependence structure between the components of $\mathbf{X}$.

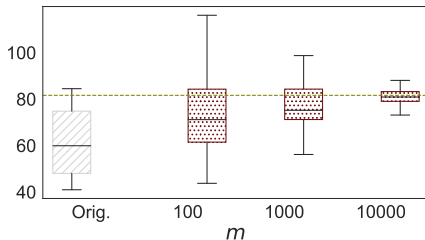
- The numerical experiments are performed on simulated data sets $\mathcal{D} \in \mathbb{R}^{1500 \times 3}$

TRMs estimation through SB simulation procedure

Estimations of the TRM on 100 original samples (grey hatched box) and on 100 extended samples (red dotted boxes). On each graph, the dashed green horizontal line corresponds to the associated theoretical value.



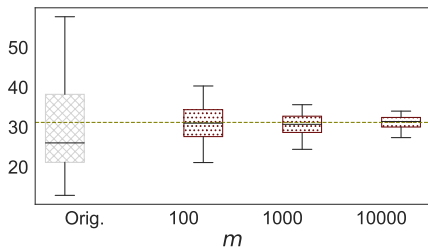
a) $ES_{0.0025}$



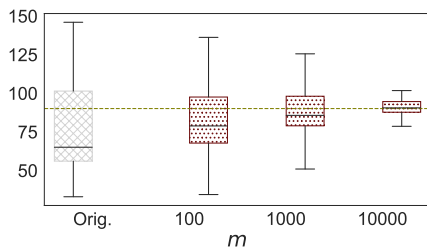
b) $ES_{0.0003}$

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Estimations of the TRM on 100 original samples (grey hatched box) and on 100 extended samples (red dotted boxes). On each graph, the dashed green horizontal line corresponds to the associated theoretical value.



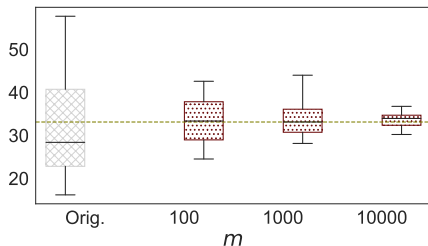
c) $MES_{0.0025}$



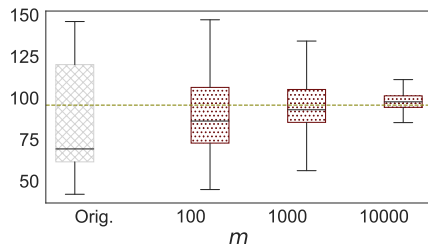
d) $MES_{0.0003}$

TRMs estimation through SB simulation procedure

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e) $DCTE_{0.0025}$



f) $DCTE_{0.0003}$

Thank you for your attention!

Reference I

- P. Artzner, F. Delbaen, J.-M. Eber, and D. Heath. Coherent measures of risk. *Mathematical finance*, 9(3):203–228, 1999.
- A. Kiriliouk, H. Rootzén, J. Segers, and J. L. Wadsworth. Peaks over thresholds modeling with multivariate generalized pareto distributions. *Technometrics*, 61(1):123–135, 2019.
- J. Legrand, P. Ailliot, P. Naveau, and N. Raillard. Joint stochastic simulation of extreme coastal and offshore significant wave heights. 2023.
- R. B. Nelsen. *An introduction to copulas*. Springer, 2006.
- H. Rootzén, J. Segers, and J. L. Wadsworth. Multivariate generalized pareto distributions: Parametrizations, representations, and properties. *Journal of Multivariate Analysis*, 165:117–131, 2018.